

Recall W8D1  
14/12/14

Prop. Methods in Comb.

Recall:  $X$  dist. Lattice,  $\mu: X \rightarrow \mathbb{R}_{\geq 0}$  is log supermodular.

FKG ineq. For increasing  $f, g: X \rightarrow \mathbb{R}_{\geq 0}$ ,  $(\sum_{x \in X} f(x)\mu(x))(\sum_{x \in X} g(x)\mu(x)) \leq (\sum_{x \in X} f(x)g(x)\mu(x))(\sum_{x \in X} \mu(x))$ .

Thm (XYZ)  $P$  ~~poset~~ finite poset,  $a_1, a_2, a_3 \in P$ . Let  $Q$  be a linear extension of  $P$  chosen UAR.

$$\text{Then } P(a_1 \leq a_3 \wedge a_2 \leq a_3) \geq P(a_1 \leq a_2)P(a_1 \leq a_3).$$

pf. Assume  $P = \{a_1, \dots, a_n\}$ . Fix  $m \gg n$  which will be specified later,  $L = \{1, \dots, m\}^n$ . Define  $\leq$  on  $L$ :

$$(x_1, \dots, x_n) \leq (y_1, \dots, y_n) \iff x_i \geq y_i \wedge \forall i \geq 2, x_i - x_1 \leq y_i - y_1.$$

claim  $(L, \leq)$  forms a lattice.

pf.  $(x_1, \dots, x_n) \vee (y_1, \dots, y_n) = (\min\{x_1, y_1\}, \min\{x_2, y_2\} + \max\{x_2 - y_2, x_1 - y_1\}, \dots)$ ,  
 $(x_1, \dots, x_n) \wedge (y_1, \dots, y_n) = (\max\{x_1, y_1\}, \max\{x_2, y_2\} + \min\{x_2 - x_1, y_2 - y_1\}, \dots)$   
 satisfy the conditions of join & meet.

$(L, \leq)$  is also distributive.

Now, define  $\mu$  the "char. function of  $P$ :  $\mu(x_1, \dots, x_n) = 1 \iff \forall i, j$  s.t.

$a_i \leq_P a_j, x_i \leq x_j$ .  $\mu$  is log-supermodular since if  $x, y$  preserve the order on  $P$ , so do  $x \vee y$  and  $x \wedge y$ .

Now, every lin. extension  $Q \supseteq P$ , if  $Q_{i_1} \leq Q_{i_2} \leq \dots \leq Q_{i_n}$ , then it is "represented" in  $L$  by some fixed num. of elements

$\bar{x}$  s.t.  $\mu(\bar{x}) = 1$ . Such a rep. may not uniquely determine  $Q$ , as equality ~~may~~ occur (but with very small prob. for large enough  $m$ ).

However,  $\bar{x} \in \mu^{-1}(1)$  which doesn't contain repeated elements uniquely determines such  $Q$ , each  $Q$  being

represented by exactly such  $\binom{n}{m}$   $\bar{x}$  values. Therefore

$\frac{1}{\binom{n}{m}} \mu|_{\mu^{-1}(1) \setminus \{\text{no repeated elements}\}}$  is the uniform measure

over linear extensions  $Q$  of  $P$ .

~~Define~~ Define  $f, g$ ,  $f(\bar{x}) = \begin{cases} 1 & x_1 \leq x_2 \\ 0 & \text{o/w} \end{cases}$   $g(\bar{x}) = \begin{cases} 1 & x_1 \leq x_3 \\ 0 & \text{o/w} \end{cases}$  By FKG

$P_L(x_1 \leq x_2) P_L(x_1 \leq x_3) \leq P_L(x_1 \leq x_2 \text{ \& } x_1 \leq x_3)$ . But  $P_L \neq$  uniform measure on lin. extensions. However,  $P(\exists i, j \text{ s.t. } x_i = x_j) \leq \frac{\binom{n}{2} m \cdot m^{n-2}}{\frac{1}{n!} m^n} \leq \frac{n! \binom{n}{2}}{m} \xrightarrow{m \rightarrow \infty} 0$ . Therefore  $P_L(A) \rightarrow P_{\#}(A)$  for the uniform measure, which gives the result.

### Concentration of Measure Inequalities

$Z$  a r.v. with  $EZ$  exists. We want to estimate  $P(Z - EZ \geq t)$  or  $P(Z - EZ \leq -t)$ , for some  $t > 0$ .

Markov:  $P(Y \geq t) \leq \frac{EY}{t}$  (can let  $Y = |Z - EZ|$ ) for  $t > 0, Y \geq 0$ .

Choose  $\varphi: \underbrace{I}_{\text{range}(Y)} \rightarrow \mathbb{R}_{\geq 0}$  increasing. Then  $P(Y \geq t) = P(\varphi(Y) \geq \varphi(t)) \leq \frac{E(\varphi(Y))}{\varphi(t)}$ .

A common choice,  $\varphi(t) = t^2$ , gives Chebyshev's inequality.

We can generalize  $\varphi(t) = t^q$  ( $q > 0$ ) which gives the moment method - but  $E\varphi(Y)$  may be difficult to calculate.

If  $Z = X_1 + \dots + X_n$  for  $X_i, X_j$  indep., we can take  $\varphi(t) = e^{\lambda t}$ .

Using Chebyshev we would get (assuming  $X_i$  i.i.d. with  $\text{Var } X_i = \sigma^2$ )

$P(|Z - EZ| \geq t) \leq \frac{n\sigma^2}{t^2}$ , which goes to 0 quad. in  $t$ , very far from the truth (exp. in  $t^2 \dots$ ).

Choose  $\lambda > 0, \varphi(t) = e^{\lambda t}$ . Then  $P(Z - EZ \geq t) \leq \frac{E(\varphi(Z - EZ))}{\varphi(t)}$ .

$E(\varphi(Z - EZ)) = \frac{\prod_{i=1}^n E \varphi(X_i)}{\prod_{i=1}^n E \varphi(X_i)}$ . So this gives a bound

$\frac{\prod_{i=1}^n E e^{\lambda X_i}}{e^{\lambda t + \lambda \sum E X_i}}$  The Cramér-Chernoff method.

Thm. Suppose  $X_1, \dots, X_n$  i.i.d. uniform on  $\{\pm 1\}$ .

Let  $S_n = X_1 + \dots + X_n$ . Then, for any  $a > 0$ ,

$P(S_n > a) = P(-S_n < a) \leq e^{-\frac{a^2}{2n}}$ . This is only good

for  $a \approx \sqrt{n} \approx \sqrt{\text{Var}(S_n)}$ , the same range of

Chebyshev, but gives much better results in this range.

pf.  $E e^{\lambda X_i} = \frac{e^\lambda + e^{-\lambda}}{2} \leq e^{\lambda^2/2}$  by comparing the Taylor series.

This gives a bound of  $\frac{e^{\lambda \sum X_i}}{e^{\lambda^2 n/2}} = e^{\lambda(S_n - \frac{\lambda n}{2})}$ . Take  $\lambda = \frac{a}{n}$  to get

$e^{-\frac{a^2}{2n}}$  as promised. In fact  $\frac{S_n}{\sqrt{n}} \rightarrow N(0,1)$  so  $P(S_n > a_n) = P\left(\frac{S_n}{\sqrt{n}} > \frac{a_n}{\sqrt{n}}\right) \xrightarrow{n \rightarrow \infty} P(N(0,1) > a) \approx \frac{e^{-\frac{a^2}{2}}}{\sqrt{2\pi}a}$  - very close to the Chernoff bound.

### Non-Symm. case

$p_1, \dots, p_n$ , so that  $P(x_i = 1 - p_i) = p_i, P(x_i = -p_i) = 1 - p_i$ , (so that  $E x_i = 0$ ). Let  $p = \frac{\sum_{i=1}^n p_i}{n} > 0$ . Then for  $X = \sum_{i=1}^n x_i$ ,

$$P(X > a) \leq \exp\left(\frac{-a^2}{2pn} + \frac{a^3}{2(pn)^2}\right), \quad P(X < -a) \leq \exp\left(\frac{-a^2}{2pn}\right).$$

both independent & indicator  
↓  
indicator

Cor ~~Under the same assumptions~~, let  $Y$  be a sum of ind. r.v.  $\forall \epsilon > 0 \exists c_\epsilon > 0$  st.  $P(|Y - EY| \geq \epsilon EY) \leq 2e^{-c_\epsilon EY}$ .

### Martingales

A sequence  $X_0, \dots, X_m$  is a martingale if each  $X_i$  is integrable and  $E[X_i | X_0, \dots, X_{i-1}] = X_{i-1}$ .

### Examples of Martingales:

- (1)  $X_i = \sum_{k=1}^i Y_k$ , where  $Y_1, \dots, Y_m$  are indep. with  $E Y_i = 0$ .
- (2) Doob Martingale -  $X$  r.v.  $\mathcal{F}_0 \subseteq \mathcal{F}_1 \subseteq \dots \subseteq \mathcal{F}_n, X_i = E[X | \mathcal{F}_i]$ .
- (3)  $[G \sim G(n, p), f: \mathcal{G}_{\text{pairs}} \rightarrow \mathbb{R}]$  Edge exposure martingale - order the  $M = \binom{[n]}{2}$  pairs of ~~v~~ v's.  $e_1, \dots, e_m$ . Define  $E[f | G \setminus \{e_1, \dots, e_i\}] = X_i$ .
- (4) Vertex exposure martingale - order the vertices  $v_1, \dots, v_n$ . Define  $X_i = E[f | G \setminus \{v_1, \dots, v_i\}]$ .

### Azuma's inequality

Let  $0 = X_0, X_1, \dots, X_m$  be a martingale s.t.  $|X_{i+1} - X_i| \leq 1$  a.s.

Then  $P(X_m > \lambda \sqrt{m}) \leq e^{-\lambda^2/2}$ .

Def. A graph func. is said to be:

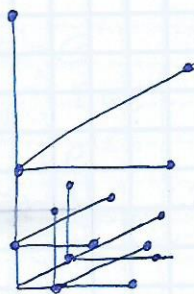
(a) edge-Lipschitz  $\Leftrightarrow |f(H) - f(H \cup \{e\})| \leq 1$ .

(b) vertex-Lipschitz  $\Leftrightarrow |f(H) - f(H \cup \{e_1, \dots, e_n\})| \leq 1$  whenever  $v \in e_1, \dots, e_n$  for some  $v$ .

Fact When  $f$  satisfies  $v$ /edge Lipschitz condition, the  $v$ /edge exposure martingale satisfies  $|X_{i+1} - X_i| \leq 1$ .

Cor. (Shamir & Spencer) For arbitrary  $n, p$ ,  $G \sim G(n, p)$ , then  $c = E[X(G)]$ . Then  $P[|X(G) - c| \geq \lambda \sqrt{n-1}] \leq 2e^{-\lambda^2/2}$ .

Azuma pf. (Sketch) Let  $Y_i = X_i - X_{i-1}$ .  $|Y_i| \leq 1$  a.s. and  $E Y_i = 0$ .  
 $E(e^{\alpha Y_i} | X_1, \dots, X_{i-1}) \leq e^{\alpha^2/2}$ , and we can just follow the pf. for Bernouli r.v.



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