

W12 D1
11/1/15

Prob. Methods in Comb.

Reminder: Yanson's inequality:

$\{B_i\}_{i \in I}$ events of the form $A_i \in \mathcal{R}$, $P(\bigcap_{i \in I} \bar{B}_i) \approx e^{-\mu}$.

$X = \sum_{i \in I} \mathbb{1}_{B_i}$, then $X \sim \text{Pois}(\mu)$.

Bonferroni Inequalities

A_1, \dots, A_n events in prob. space Ω . Given $I \subseteq [n]$

let $A_I = \bigcap_{i \in I} A_i$. Given $\epsilon \in \{0, 1\}^n$, $A_\epsilon = \bigcap_{i \in [n]} A_i^{\epsilon_i}$ where $A_i^0 = \bar{A}_i, A_i^1 = A_i$.

$P(A_{\epsilon_0, \dots, \epsilon_n}) = \sum_{I \subseteq [n]} \sum_{\Pi \subseteq I} (-1)^{|\Pi|} P(A_\Pi)$. Bonferroni - $\forall s \geq 0$,

$\sum_{I=0}^{2s+1} (-1)^I P(A_\Pi) \leq P(A_{\epsilon_0, \dots, \epsilon_n})$
inc. ex. $\sum_{I=0}^s (-1)^I P(A_\Pi)$

Brun's Sieve

$n \rightarrow \infty$, $\Omega(n)$ prob. space & $A_1(n), \dots, A_m(n)$ events in it.

$X(n) = \sum_{i=1}^m \mathbb{1}_{A_i(n)}$, $S^{(r)} = S^{(r)}(n) = \sum_{I \subseteq [m], |I|=r} P(A_I(n))$. Suppose $\exists p$ const.

s.t. $\forall r \geq 1$, $\lim_{n \rightarrow \infty} S^{(r)}(n) = \frac{\mu^r}{r!}$. Then for every fixed t ,

$\lim_{n \rightarrow \infty} P(X=t) = \frac{\mu^t}{t!} e^{-\mu} = P(\text{Pois}(\mu)=t)$.
in particular $E[X] \rightarrow \mu$

pf. for $t=0$, fix $\delta > 0$. First choose S_0 s.t. $\forall s > S_0$, $|e^{-\mu} - \sum_{r=0}^s \frac{\mu^r}{r!} (-1)^r| < \delta/2$.

Choose n_0 s.t. $\forall n > n_0$, $|S^{(r)}(n) - \frac{\mu^r}{r!}| < \frac{\delta}{2(s_0+2)}$ for

all $r \leq 2s_0+1$. By Bonferroni, $\sum_{r=0}^{2s_0+1} (-1)^r \frac{\mu^r}{r!} - \delta/2 \leq \sum_{r=0}^{2s_0+1} S^{(r)}(n) \leq P(X=0) \leq$

$\sum_{r=0}^{2s_0+1} (-1)^r S^{(r)}(n) \leq \sum_{r=0}^{2s_0+1} \frac{\mu^r}{r!} + \delta \leq e^{-\mu} + \delta$. For $t \neq 0$ fix $T \subseteq I$ of size t .

$P(\text{Only } A_i \text{ with } i \in T \text{ hold})$ can be estimated using the prev.

methods using $X_T = \# \text{ events } \& T \text{ which hold}$, and summing over $|T|=t$ we

get $\frac{\mu^t}{t!} e^{-\mu} = \sum_{r=0}^{\infty} \binom{r+t}{t} \frac{\mu^{r+t}}{(r+t)!}$

Example: Threshold func. for $\forall v, x \in \text{Copy of } K_3$ -

Let $p = p(n)$, $\mu = \mu(n)$ s.t. $e^{-\mu} = c/n$, $\mu = \binom{n-1}{2} c^3$ (so

$\mu \approx \log n$, $p \approx n^{-2/3} (\log n)^{1/3}$). Then $P(G(n, p) \text{ has the}$

property) $= e^{-c}$.

pf. $G \sim G(n, p)$. $B_{xyz} = \{xyz \text{ induce a triangle}\}$. $A_x = \bigcap_{y, z \neq x} B_{xyz}$.

Apply Yanson with $\bar{\mu} = E[\# B_{xyz} \text{ that hold}] = p^3 \binom{n-1}{2} = \mu$ and

$\lambda = n^{-2/3 + o(1)} = o(1)$, so $P(A_x) \leq e^{-\mu + o(1)} = (1 + o(1)) e^{-\mu}$, and by

Harri's inequality /FKG $P(A_x) \geq (1-p^3)^{\binom{n-1}{2}} \geq (1-o(1))e^{-\gamma}$. Now put $X = \sum_{x \in V} \mathbb{1}_{A_x}$. $E[X] = S^{(2)} = (1+o(1))e^{-\gamma} \cdot n = (1+o(1))c$. We want $P(X=0) \rightarrow e^{-c}$ so

it's enough $S^{(2)} \rightarrow \frac{c^2}{r!}$. Let $\Pi = r$, $A_r = \bigcap_{x \in T} A_x = \bigcap_{S: |S|=r} B_S$, $\mu_r = [r \binom{n-1}{2} - o(n)]p^3$

$\Delta = O(p^5 \cdot r \cdot n^3) = o(1)$, so $P(A_r) = \prod_{S: |S|=r} P(B_S) = e^{-\mu_r} (1+o(1))$. So

$$S^{(r)} = \binom{n}{r} e^{-\mu_r} (1+o(1)) = \frac{n^r e^{-\mu_r}}{r!} = \frac{(ne^{-\mu_r})^r}{r!} \stackrel{\text{by Yanson-FKG}}{=} \frac{c^r}{r!} (1+o(1)).$$

Large Deviation Ineq. for sums of "Almost Indep." r.v.s

Ω fin. $R \subseteq \Omega$ randomly $P(x \in R) = p_x$ indep. $B_i = \mathbb{1}_{A_i \in R}$.

$i \sim j \Leftrightarrow i \neq j, A_i \cap A_j \neq \emptyset$. We call $J \subseteq I$ a disjoint family if J is

\sim -indep. \Rightarrow Maximal disjoint family if J can't be extended.

Thm. \forall integers, $\mu = \sum_{i \in I} P(B_i)$, $P(\exists J | J|=s \text{ disj. family}) \leq \frac{\mu^s}{s!} \leq (\frac{e\mu}{s})^s$. In

particular, if $s = (e+\delta)\mu$, this $P \leq (1-\rho(\delta))^\mu$.

pf. $\mathcal{A} = \{J \subseteq I \text{ } \sim\text{-indep.}\}$. $\text{ord} \{j_1, \dots, j_s\} \in I^s \mid \{j_1, \dots, j_s\} \sim\text{-ind.}\}$

$$P(\exists J | J|=s \text{ disj.}) \leq \sum_{J \in \mathcal{A}} P(\bigcap_{j \in J} B_j) = \sum_{J \in \mathcal{A}} \prod_{j \in J} P(B_j) = \frac{1}{s!} \sum_{J \in \mathcal{A}} P(B_{j_1}) \dots P(B_{j_s}) \leq$$

$$\frac{1}{s!} \sum_{J \in I^s} P(B_{j_1}) \dots P(B_{j_s}) = \frac{1}{s!} \left(\sum_{j \in I} P(B_j) \right)^s = \frac{\mu^s}{s!}.$$

Illustration - off-diag. Ramsey nums. (Krivelevich)

Fix $l \geq 3$. $R(l, k) = \max \{n \mid \exists G \text{ on } n \text{ vxs. } G \not\cong K_l, \omega(G) < k\}$.

prop. $R(l, k) \geq c_l \left(\frac{k}{\log k} \right)^{\frac{l-1}{2}}$ for some $c_l > 0$.

For n above $\mu = c' n^{\frac{2}{l-1}}$, $n^l p^{\binom{l}{2}} = \Theta(n^2 p^{\frac{l}{2}})$. For $S \subseteq V(G)$

$X_S = e(G[S])$, $Y_S = \#$ copies in $G[S]$ with ≥ 2 vxs. in S , $Z_S = \#$ max edge disjoint such copies. $A_S = "X_S \geq \binom{l}{2} Z_S"$.

If $\bigcap_{|S|=k} A_S$ holds, $G(n, p)$ witnesses $R(k, l) \geq n$. Let H be a family in

the def. of Z_S . Delete from G edges in those cliques-

$e_G(S) \geq X_S - \binom{l}{2} Z_S > 0$ for any $S \in \binom{[n]}{k}$ in the resulting graph, and

it's also K_l -free. Now, for fixed $S \in \binom{[n]}{k}$, $P(\overline{A}_S) \leq P(X_S \leq \frac{E X_S}{2}) +$

$P(\binom{l}{2} Z_S \geq \frac{E X_S}{2})$, $E[X_S] = \binom{k}{2} p$, so $\leq e^{-\binom{l}{2} p/8}$. For 2nd part,

$E(Z_S) \leq E(Y_S) = \mu_S \leq \binom{k}{2} \binom{n-2}{l-2} p^{\binom{l}{2}}$, $\mu_S \geq \binom{k}{2} \binom{n-k}{l-2} p^{\binom{l}{2}}$ but $k = o(n)$ so

they're both $\Theta(\binom{k}{2} n^{l-2} p^{\binom{l}{2}})$. By lemma $P(Z_S \geq 5\mu_S) \leq (e/5)^{5\mu_S}$. So

$\frac{c_0 \leq \mathbb{E}[X_s]}{\mu_s} \leq \frac{\binom{k}{s} p}{c_0 \binom{k-2}{s-2} p^2} = \frac{1}{c_0 (c')^{s-1}}$ for small enough $c_0 > 0$. For

small enough c' , $10(\frac{1}{2}) \leq \frac{\mathbb{E}[X_s]}{\mu_s} \leq D = \Theta(1)$. With this choice of c' , $\mathbb{P}(\binom{l}{2} Z_s \geq \mathbb{E} X_s / 2) \leq \mathbb{P}(Z_s \geq 5\mu_s) \leq (e/5)^{\frac{\mathbb{E}[X_s]}{\sigma}} = e^{-\tilde{c} \binom{k}{2} p}$.

Now we know $\mathbb{P}(A_s) \leq 2e^{-\tilde{c} \binom{k}{2} p}$. So, $\mathbb{P}(\bigwedge_{|S|=k} A_S) \geq 1 - \sum_{|S|=k} \mathbb{P}(A_S) \geq 1 - \binom{n}{k} e^{-\tilde{c} \binom{k}{2} p} \geq 1 - (\frac{en}{k}) e^{-\tilde{c} \frac{k-1}{2} p}$. (2) follows from $pk \geq \frac{2l}{\varepsilon} \log k \Leftrightarrow p \geq \frac{2l}{\varepsilon k} \log k = \frac{2l}{\varepsilon k} \left(\frac{n}{k}\right)^{-\frac{2}{k}} = \frac{2l}{\varepsilon c^{\frac{2}{k}}} \frac{p}{c}$ and we can just choose c sufficiently small.

† Poisson concentration of measure - $\mathbb{P}(P \leq (1-\varepsilon)\mu) \leq e^{-\varepsilon^2 \mu / 2}$, $\mathbb{P}(P \geq (1+\varepsilon)\mu) \leq [e^{-\varepsilon} (1+\varepsilon)]^\mu$. Define $s \geq 0$, $\mu_s = \min_{J \subseteq I, |J|=s} \sum_{i \in J} \mathbb{P}(B_i)$.

Letting $\nu = \max_{i \in I} \sum_{j \in I} \mathbb{P}(B_j)$, then $\mu_s \geq \nu - s\nu$.

Lemma $\mathbb{P}(\max_{\text{disj. family of sizes } s} \sum_{i \in J} \mathbb{P}(B_i)) \leq \frac{\mu^s}{s!} e^{-\mu_s + 1/2}$.